# Multi-innovation gradient identification for input nonlinear state space systems $\star$

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**Abstract:** This paper presents a multi-innovation stochastic gradient algorithm for an input nonlinear state space system by deriving the identification model of the system and by decomposing the model into two sub-models. The basic idea is to design a state observer to estimate the unmeasurable states and to estimate interactively the unknown parameters of two subsystems by using the hierarchical identification principle. The simulation results show that the proposed algorithm is efficient.

Keywords: Recursive identification, Gradient search, Multi-innovation identification, State space model

## 1. INTRODUCTION

Nonlinearities are often encountered in fault detection and system identification (Jiang & Chowdhury, 2005; Zhang, et al., 2014; Wang & Ding, 2014). The nonlinear system modeling has been widely used in the industrial processes, such as the microbial batch fermentation and pH process (Chen, et al., 2014; Liu, 2013). Much effort was devoted to the identification of the block-oriented nonlinear systems, which include the input nonlinear systems, i.e., a static nonlinear block followed by a linear dynamic one, and the output nonlinear systems as the special cases (Chai, Loxton, Teo, & Yang, 2013; Enqvist & Ljung, 2005; Vörös, 2007). For example, Yu, Zhang, and Xie (2014) studied the parameter estimation problem for Hammerstein systems using a blind identification based deterministic estimation algorithm. Li and Wen (2013) transformed the known basis functions into the orthogonal basis functions and proposed a normalized iterative algorithm for Hammerstein systems by applying a fixed-point iteration technique.

Some algorithms have been reported for the nonlinear systems, such as the over-parametrization methods (Bai, 1998; Liu & Bai, 2007), the least squares methods (Krstic, 2009; Vörös, 2005), the iterative methods (Bai & Li, 2004; Vörös, 1999; Wang & Tang, 2014) and the maximum-likelihood methods (Hagenblad, Ljung, & Wills, 2008; Li, 2013). The over-parametrization methods transform the output of the original nonlinear system into a linear function on the parameter space so that any linear estimation algorithms can be applied (Ding & Chen, 2005). However, the resulting cross-product increases the dimensions of the

identification parameters and leads to a heavy computational burden. The hierarchical identification algorithm is to decompose a system into two subsystems with small dimensions, and to identify the parameters of each subsystem respectively, thus the computational efficiency can be improved (Ding, 2013). Raja and Chaudhary (2014) presented a two-stage fractional least mean squares adaptive algorithm for controlled autoregressive moving average systems. Ding and Duan (2013) derived a two-stage recursive least squares algorithm and a two-stage gradient algorithm for Box-Jenkins models.

State space models can describe the dynamic of the systems and play a key role in the controller design (Dhawan & Kar, 2011; Wu, Wang, & Li, 2012), signal filtering (Levy & Nikoukhah, 2013) and system analysis (Hmida, Khémiri, Ragot, & Gossa, 2012; Mulders, Schoukens, & Vanbeylen, 2013; Pence, Fathy, & Stein, 2011). The identification of the state space models may involve not only the estimation of the unknown model parameters, but also the unmeasurable system states. The combined state and parameter estimation based recursive identification is important since it can be implemented with advanced control strategies or monitoring process behavior to improve the online control performance of the key quality variables. The maximum likelihood method or the expectation-maximization method have been reported for estimating the state and parameter simultaneously (Mader, Linke, et al., 2014). Tulsyan, Huang, Gopaluni, and Forbes (2013) presented a Bayesian approach for nonlinear state space models with an adaptive sequential-importance-resampling filter. Schön, Wills, and Ninness (2011) proposed an expectationmaximization algorithm for nonlinear dynamic systems in state-space form using a particle smoother.

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The importance of the canonical forms of the state space models is widely admitted. Mercère and Bako (2011) presented the canonical form of the multivariable systems and discussed its application in the subspace identification. Based on the observability canonical form, Ding (2014) derived a combined state and least squares algorithm for state space systems by means of the Kalman filter. Recently, for the nonlinear state space systems with onetime delay, Gu and Ding (2014) investigated the gradient based and the least squares based iterative algorithms. Wang, Ding, and Liu (2014) studied the least squares based parameter and state estimation algorithm for the input nonlinear state space systems. This paper develops a highly accurate recursive identification algorithm for input nonlinear state space systems using the hierarchical identification and the multi-innovation identification theory. The multi-innovation identification algorithm is introduced to improve the computational efficiency of the stochastic gradient methods by utilizing current innovations and past innovations repeatedly.

The rest of this paper is organized as follows. Section 2 derives the identification model for input nonlinear state space systems. Section 3 develops a multi-innovation stochastic gradient (MISG) algorithm to estimate the system parameters and states jointly. Section 4 provides an illustrative example to show the effectiveness of the proposed algorithms. Finally, concluding remarks are given in Section 5.

### 2. PROBLEM DESCRIPTION AND SYSTEM MODEL

Consider the following input nonlinear state space systems,

$$\boldsymbol{x}(t+1) = \boldsymbol{A}\boldsymbol{x}(t) + \boldsymbol{b}\bar{\boldsymbol{u}}(t), \qquad (1)$$

$$\bar{u}(t) = g(u(t)), \tag{2}$$

$$y(t) = \boldsymbol{c}^{\mathrm{T}}\boldsymbol{x}(t) + w(t), \qquad (3)$$

$$w(t) = H(z)v(t), \tag{4}$$

where  $\boldsymbol{x}(t) := [x_1(t), x_2(t), \cdots, x_n(t)]^{\mathrm{T}} \in \mathbb{R}^n$  is the state vector,  $\bar{u}(t) \in \mathbb{R}$  is the output of the nonlinear part and unmeasurable,  $u(t) \in \mathbb{R}$  is the system input,  $y(t) \in \mathbb{R}$  is the system output,  $v(t) \in \mathbb{R}$  is a random noise with zero mean,  $\boldsymbol{A} \in \mathbb{R}^{n \times n}$ ,  $\boldsymbol{b} \in \mathbb{R}^n$  and  $\boldsymbol{c} \in \mathbb{R}^n$  are the parameter matrix and vectors of the system,

$$\boldsymbol{A} := \begin{bmatrix} 0 & 1 & 0 & \cdots & 0 \\ 0 & 0 & 1 & \ddots & \vdots \\ \vdots & \vdots & \ddots & \ddots & 0 \\ 0 & 0 & \cdots & 0 & 1 \\ a_1 & a_2 & a_3 & \cdots & a_n \end{bmatrix} \in \mathbb{R}^{n \times n},$$
$$\boldsymbol{b} := [b_1, b_2, \cdots, b_n]^{\mathrm{T}} \in \mathbb{R}^n,$$
$$\boldsymbol{c} := [1, 0, \cdots, 0]^{\mathrm{T}} \in \mathbb{R}^n.$$

The disturbance  $w(t) \in \mathbb{R}$  is a process noise, H(z) is a finite impulse response model:

$$H(z) = 1 + h_1 z^{-1} + h_2 z^{-2} + \dots + h_{n_h} z^{-n_h}$$

The nonlinear output  $\bar{u}(t)$  is a linear combination of a known basis with coefficients  $\boldsymbol{\gamma} := (\gamma_1, \gamma_2, \cdots, \gamma_m)^{\mathrm{T}} \in \mathbb{R}^m$ , and can be expressed as

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$$\bar{u}(t) = g(u(t)) = \sum_{j=1}^{m} \gamma_j g_j(u(t)).$$
 (5)

From (1), we have

$$x_{1}(t) = a_{1}x_{1}(t-n) + a_{2}x_{2}(t-n) + \dots + a_{n}x_{n}(t-n) + b_{1}\bar{u}(t-1) + b_{2}\bar{u}(t-2) + \dots + b_{n}\bar{u}(t-n) = \boldsymbol{x}^{\mathrm{T}}(t-n)\boldsymbol{a} + \sum_{i=1}^{n} b_{i}\sum_{j=1}^{m}\gamma_{j}g_{j}(u(t-i)) = \boldsymbol{x}^{\mathrm{T}}(t-n)\boldsymbol{a} + \boldsymbol{b}^{\mathrm{T}}\boldsymbol{G}(t)\boldsymbol{\gamma},$$
(6)

where

$$\boldsymbol{G}(t) := \begin{bmatrix} g_1(u(t-1)) & g_2(u(t-1)) & \cdots & g_m(u(t-1)) \\ g_1(u(t-2)) & g_2(u(t-2)) & \cdots & g_m(u(t-2)) \\ \vdots & \vdots & & \vdots \\ g_1(u(t-n)) & g_2(u(t-n)) & \cdots & g_m(u(t-n)) \end{bmatrix} \\ \in \mathbb{R}^{n \times m}.$$
(7)

From (4), we have

$$w(t) = \sum_{i=1}^{n_h} h_i v(t-i) + v(t)$$
$$= \boldsymbol{\varphi}^{\mathrm{T}}(t) \boldsymbol{h} + v(t), \qquad (8)$$

where

$$\boldsymbol{\varphi}(t) := [v(t-1), v(t-2), \cdots, v(t-n_h)]^{\mathrm{T}} \in \mathbb{R}^{n_h},$$
$$\boldsymbol{h} := [h_1, h_2, \cdots, h_{n_h}]^{\mathrm{T}} \in \mathbb{R}^{n_h}.$$

Inserting (6) and (8) into (3) yields

$$y(t) = x_1(t) + w(t)$$

$$= \boldsymbol{x}^{\mathrm{T}}(t-n)\boldsymbol{a} + \boldsymbol{b}^{\mathrm{T}}\boldsymbol{G}(t)\boldsymbol{\gamma} + \boldsymbol{\varphi}^{\mathrm{T}}(t)\boldsymbol{h} + v(t).$$
(10)

Define the intermediate variable

$$y_1(t) := y(t) - \boldsymbol{x}^{\mathrm{T}}(t-n)\boldsymbol{a} - \boldsymbol{\varphi}^{\mathrm{T}}(t)\boldsymbol{h}$$
  
and the information vectors

$$\begin{split} \boldsymbol{\eta}(\boldsymbol{\gamma},t) &:= \begin{bmatrix} \boldsymbol{x}(t-n) \\ \boldsymbol{G}(t)\boldsymbol{\gamma} \\ \boldsymbol{\varphi}(t) \end{bmatrix} \in \mathbb{R}^{2n+n_h}, \\ \boldsymbol{\xi}(\boldsymbol{b},t) &:= \boldsymbol{G}^{\mathrm{T}}(t)\boldsymbol{b} \in \mathbb{R}^{m}. \end{split}$$

The model in (10) can be decomposed into the following two identification sub-models,

$$S_{1}: y(t) = [\boldsymbol{x}^{\mathrm{T}}(t-n), \boldsymbol{\gamma}^{\mathrm{T}}\boldsymbol{G}^{\mathrm{T}}(t), \boldsymbol{\varphi}^{\mathrm{T}}(t)]\boldsymbol{\vartheta} + v(t)$$
  
=  $\boldsymbol{\eta}^{\mathrm{T}}(\boldsymbol{\gamma}, t)\boldsymbol{\vartheta} + v(t),$  (11)  
$$S_{2}: y_{1}(t) = \boldsymbol{b}^{\mathrm{T}}\boldsymbol{G}(t)\boldsymbol{\gamma} + v(t)$$

$$=\boldsymbol{\xi}^{\mathrm{T}}(\boldsymbol{b},t)\boldsymbol{\gamma}+v(t). \tag{12}$$

The parameter vectors  $\boldsymbol{a}, \boldsymbol{b}, \boldsymbol{\gamma}$  and  $\boldsymbol{h}$  to be estimated are included in sub-models (11) and (12).

## 3. THE MULTI-INNOVATION STOCHASTIC GRADIENT ALGORITHM

Consider p data from j = t - p + 1 to j = t. Define the cost functions

$$\begin{split} J_1(\boldsymbol{\vartheta}) &:= \sum_{j=0}^{p-1} [y(t-j) - \boldsymbol{\eta}^{\mathrm{T}}(\boldsymbol{\gamma}, t-j)\boldsymbol{\vartheta}]^2, \\ J_2(\boldsymbol{\gamma}) &:= \sum_{j=0}^{p-1} [y_1(t-j) - \boldsymbol{\xi}^{\mathrm{T}}(\boldsymbol{b}, t-j)\boldsymbol{\gamma}]^2 \\ &= \sum_{j=0}^{p-1} [y(t-j) - \boldsymbol{\xi}^{\mathrm{T}}(\boldsymbol{b}, t-j)\boldsymbol{\gamma} \\ &- \boldsymbol{x}^{\mathrm{T}}(t-n-j)\hat{\boldsymbol{a}}(t) - \boldsymbol{\varphi}^{\mathrm{T}}(t-j)\hat{\boldsymbol{h}}(t)]^2 \end{split}$$

Note that  $\gamma$  in  $J_1(\vartheta)$  and  $\boldsymbol{b}$  in  $J_2(\gamma)$  are unknown and can be replaced by their corresponding estimates  $\hat{\gamma}(t-1)$  and  $\hat{\boldsymbol{b}}(t)$ . Using the gradient search and minimizing  $J_1(\vartheta)$  and  $J_2(\gamma)$  give

$$\hat{\vartheta}(t) = \hat{\vartheta}(t-1) + \frac{1}{r_1(t)} \sum_{j=0}^{p-1} \hat{\eta}(\hat{\gamma}(t-1), t-j) \\ \times [y(t-j) - \hat{\eta}^{\mathrm{T}}(\hat{\gamma}(t-1), t-j)\hat{\vartheta}(t-1)],$$
(13)

$$r_1(t) = r_1(t-1) + \|\hat{\boldsymbol{\eta}}(\hat{\boldsymbol{\gamma}}(t-1), t)\|^2, \tag{14}$$

$$\hat{\boldsymbol{\gamma}}(t) = \hat{\boldsymbol{\gamma}}(t-1) + \frac{1}{r_2(t)} \sum_{j=0}^{p-1} \boldsymbol{\xi}(\hat{\boldsymbol{b}}(t), t-j) \\ \times [y(t-j) - \boldsymbol{\xi}^{\mathrm{T}}(\hat{\boldsymbol{b}}(t), t-j)\hat{\boldsymbol{\gamma}}(t-1) \\ -\hat{\boldsymbol{x}}^{\mathrm{T}}(t-n-j)\hat{\boldsymbol{a}}(t) - \hat{\boldsymbol{\varphi}}^{\mathrm{T}}(t-j)\hat{\boldsymbol{h}}(t)], \qquad (15)$$

$$r_2(t) = r_2(t-1) + \|\boldsymbol{\xi}(\hat{\boldsymbol{b}}(t), t)\|^2, \qquad (16)$$

where p represents the innovation length. From (8) and (3), we have

$$\begin{aligned} \boldsymbol{v}(t) &= \boldsymbol{w}(t) - \boldsymbol{\varphi}^{\mathrm{T}}(t)\boldsymbol{h} \\ &= \boldsymbol{y}(t) - \boldsymbol{x}_{1}(t) - \boldsymbol{\varphi}^{\mathrm{T}}(t)\boldsymbol{h} \end{aligned}$$

Let  $\hat{\boldsymbol{\varphi}}(t)$  be the estimate of  $\boldsymbol{\varphi}(t)$ ,

$$\hat{\boldsymbol{\varphi}}(t) := [\hat{v}(t-1), \hat{v}(t-2), \cdots, \hat{v}(t-n_h)]^{\mathrm{T}} \in \mathbb{R}^{n_h}.$$
 (17)  
Define

$$\hat{\boldsymbol{x}}(t) := [\hat{x}_1(t), \hat{x}_2(t), \cdots, \hat{x}_n(t)]^{\mathrm{T}} \in \mathbb{R}^n.$$
(18)

Replacing  $x_1(t)$ ,  $\varphi(t)$  and h with  $\hat{x}_1(t)$ ,  $\hat{\varphi}(t)$  and  $\hat{h}(t)$ , the estimate of v(t) can be computed through

$$\hat{v}(t) = y(t) - \hat{x}_1(t) - \hat{\boldsymbol{\varphi}}^{\mathrm{T}}(t)\hat{\boldsymbol{h}}(t).$$
(19)

Define the stacked output vector  $\boldsymbol{Y}(p,t)$  and the stacked information matrices  $\boldsymbol{\Gamma}(p,t)$ ,  $\boldsymbol{\Xi}(p,t)$ ,  $\boldsymbol{X}(p,t)$ ,  $\boldsymbol{\Psi}(p,t)$  and  $\boldsymbol{\Phi}(p,t)$  as

$$\boldsymbol{Y}(p,t) := [y(t), y(t-1), \cdots, y(t-p+1)]^{\mathrm{T}} \in \mathbb{R}^{p}, \quad (20)$$
$$\boldsymbol{\Gamma}(p,t) := [\hat{\boldsymbol{p}}(\hat{\boldsymbol{\gamma}}(t-1), t), \hat{\boldsymbol{p}}(\hat{\boldsymbol{\gamma}}(t-1), t-1), \cdots, p(t-p+1)]^{\mathrm{T}}$$

$$\hat{\boldsymbol{\eta}}(\hat{\boldsymbol{\gamma}}(t-1), t-p+1)] \in \mathbb{R}^{(2n+n_h) \times p}, \qquad (21)$$

$$\boldsymbol{\Xi}(p,t) := [\boldsymbol{\xi}(\hat{\boldsymbol{b}}(t),t), \boldsymbol{\xi}(\hat{\boldsymbol{b}}(t),t-1), \cdots, \\ \boldsymbol{\xi}(\hat{\boldsymbol{b}}(t),t-p+1)] \in \mathbb{R}^{m \times p},$$
(22)

$$\boldsymbol{X}(p,t) := [\hat{\boldsymbol{x}}(t-n), \hat{\boldsymbol{x}}(t-n-1), \cdots, \hat{\boldsymbol{x}}(t-n-p+1)] \\ \in \mathbb{R}^{n \times p},$$
(23)

$$\mathbf{\Phi}(p,t) := [\hat{\boldsymbol{\varphi}}(t), \hat{\boldsymbol{\varphi}}(t-1), \cdots, \hat{\boldsymbol{\varphi}}(t-p+1)] \in \mathbb{R}^{n_h \times p}.$$

Introduce two innovation vectors

$$\begin{split} \boldsymbol{E}_{1}(\boldsymbol{p},t) &:= \boldsymbol{Y}(\boldsymbol{p},t) - \boldsymbol{\Gamma}^{\mathrm{T}}(\boldsymbol{p},t) \hat{\boldsymbol{\vartheta}}(t-1) \in \mathbb{R}^{p}, \\ \boldsymbol{E}_{2}(\boldsymbol{p},t) &:= \boldsymbol{Y}(\boldsymbol{p},t) - \boldsymbol{\Xi}^{\mathrm{T}}(\boldsymbol{p},t) \hat{\boldsymbol{\gamma}}(t-1) \\ &- \boldsymbol{X}^{\mathrm{T}}(\boldsymbol{p},t) \hat{\boldsymbol{a}}(t) - \boldsymbol{\Phi}^{\mathrm{T}}(\boldsymbol{p},t) \hat{\boldsymbol{h}}(t) \in \mathbb{R}^{p} \end{split}$$

Then Equations (13) and (15) can equivalently be rewritten as

$$\hat{\boldsymbol{\vartheta}}(t) = \hat{\boldsymbol{\vartheta}}(t-1) + \frac{\boldsymbol{\Gamma}(p,t)}{r_1(t)} \boldsymbol{E}_1(p,t),$$
$$\hat{\boldsymbol{\gamma}}(t) = \hat{\boldsymbol{\gamma}}(t-1) + \frac{\boldsymbol{\Xi}(p,t)}{r_2(t)} \boldsymbol{E}_2(p,t).$$

The unknown state  $\boldsymbol{x}(t)$  in the information vector  $\boldsymbol{\eta}(\hat{\boldsymbol{\gamma}}(t-1),t)$  can be replaced with the state observer  $\hat{\boldsymbol{x}}(t)$  in the preceding section, and the unknown v(t-i) are replaced with its estimate  $\hat{v}(t-i)$ . Thus we can summarize the multi-innovation gradient identification algorithm:

$$\hat{\boldsymbol{\vartheta}}(t) = \hat{\boldsymbol{\vartheta}}(t-1) + \frac{\boldsymbol{\Gamma}(p,t)}{r_1(t)} \boldsymbol{E}_1(p,t), \qquad (25)$$

$$r_1(t) = r_1(t-1) + \|\hat{\boldsymbol{\eta}}(\hat{\boldsymbol{\gamma}}(t-1), t)\|^2, \qquad (26)$$

$$\hat{\gamma}(t) = \hat{\gamma}(t-1) + \frac{\Xi(p,t)}{r_2(t)} E_2(p,t),$$
 (27)

$$r_{2}(t) = r_{2}(t-1) + \|\boldsymbol{\xi}(\hat{\boldsymbol{b}}(t),t)\|^{2}, \qquad (28)$$

$$\boldsymbol{E}_{1}(\boldsymbol{p},t) = \boldsymbol{Y}(\boldsymbol{p},t) - \boldsymbol{\Gamma}^{\mathrm{T}}(\boldsymbol{p},t)\boldsymbol{\vartheta}(t-1), \quad (29)$$
$$\boldsymbol{E}_{2}(\boldsymbol{p},t) = \boldsymbol{Y}(\boldsymbol{p},t) - \boldsymbol{\Xi}^{\mathrm{T}}(\boldsymbol{p},t)\hat{\boldsymbol{\gamma}}(t-1)$$

$$-\boldsymbol{X}^{\mathrm{T}}(\boldsymbol{p},t)\hat{\boldsymbol{a}}(t) - \boldsymbol{\Phi}^{\mathrm{T}}(\boldsymbol{p},t)\hat{\boldsymbol{h}}(t), \qquad (30)$$

$$\boldsymbol{\xi}(\hat{\boldsymbol{b}}(t), t) = \boldsymbol{G}^{\mathrm{T}}(t)\hat{\boldsymbol{b}}(t), \tag{31}$$

$$\hat{\boldsymbol{\eta}}(\hat{\boldsymbol{\gamma}}(t-1),t) = \begin{bmatrix} \hat{\boldsymbol{x}}(t-n) \\ \boldsymbol{G}(t)\hat{\boldsymbol{\gamma}}(t-1) \\ \hat{\boldsymbol{\varphi}}(t) \end{bmatrix}, \hat{\boldsymbol{\vartheta}}(t) = \begin{bmatrix} \boldsymbol{a}(t) \\ \hat{\boldsymbol{b}}(t) \\ \hat{\boldsymbol{b}}(t) \\ \hat{\boldsymbol{h}}(t) \end{bmatrix}, (32)$$

$$\hat{\boldsymbol{\gamma}}(t) = [\hat{\gamma}_1(t), \hat{\gamma}_2(t), \cdots, \hat{\gamma}_m(t)]^{\mathrm{T}}.$$
(33)

Based on the Kalman filtering theory, the following state observer algorithm can be used for estimating the state vector  $\boldsymbol{x}(t)$ ,

$$\hat{\boldsymbol{x}}(t+1) = \hat{\boldsymbol{A}}(t)\hat{\boldsymbol{x}}(t) + \hat{\boldsymbol{b}}(t)\hat{\bar{\boldsymbol{u}}}(t), \qquad (34)$$

$$\hat{\boldsymbol{A}}(t) = \begin{bmatrix} 0 & 1 & 0 & \cdots & 0 \\ 0 & 0 & 1 & \ddots & \vdots \\ \vdots & \vdots & \ddots & \ddots & 0 \\ 0 & 0 & \cdots & 0 & 1 \\ \hat{a}_1(t) & \hat{a}_2(t) & \hat{a}_3(t) & \cdots & \hat{a}_n(t) \end{bmatrix}, (35)$$

$$\hat{\boldsymbol{b}}(t) = [\hat{b}_1(t), \hat{b}_2(t), \cdots, \hat{b}_n(t)]^{\mathrm{T}},$$
(36)

$$\hat{\bar{u}}(t) = \sum_{i=1}^{m} \hat{\gamma}_i(t) g_i(u(t)).$$
(37)

Equations (17)-(37) form the multi-innovation stochastic gradient algorithm for Hammerstein state space systems.

#### 4. EXAMPLE

Consider the following Hammerstein state space system,

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$$\begin{aligned} \boldsymbol{x}(t+1) &= \begin{bmatrix} 0 & 1\\ 0.84 & 0.13 \end{bmatrix} \boldsymbol{x}(t) + \begin{bmatrix} 0.71\\ 0.7042 \end{bmatrix} \bar{u}(t), \\ y(t) &= [1, 0] \boldsymbol{x}(t) + H(z) v(t), \\ \bar{u}(t) &= 0.05 u(t) + 0.09 u^2(t), \\ H(z) &= 1 + 0.43 z^{-1}. \end{aligned}$$

The parameter vector to be estimated is

$$\Theta := [a_1, a_2, b_1, b_2, h_1, \gamma_1, \gamma_2]^{\mathrm{T}}$$
  
= [0.84, 0.13, 0.71, 0.7042, 0.430, 0.05, 0.09]^{\mathrm{T}},

where  $\|\boldsymbol{b}\|^2 = b_1^2 + b_2^2 = 1$ . In simulation, the input  $\{u(t)\}$  is taken as an uncorrelated stochastic signal sequence with zero mean and unit variance, and  $\{v(t)\}$  as a white noise sequence with zero mean and variance  $\sigma^2 = 0.30^2$ . Take the data length L = 3000 and use the MISG algorithm to estimate the parameters of this system. The parameter estimates with different innovation lengths p = 1, p = 3 and p = 8 are shown in Tables 1–3, and the estimation errors  $\delta := \|\hat{\boldsymbol{\Theta}}(t) - \boldsymbol{\Theta}\| / \|\boldsymbol{\Theta}\|$  versus t are shown in Figure 1.

From Tables 1–3 and Figure 1, we can draw the following conclusions.

- The estimation errors given by the stochastic gradient algorithm and the MISG algorithm become smaller with increasing t see Tables 1–3.
- For the MISG algorithm, a larger innovation length p leads to a faster convergence rate see Figure 1.

## 5. CONCLUSIONS

A multi-innovation stochastic gradient identification algorithm is presented for input nonlinear state space systems. The simulation test indicates that increasing the innovation length can improve the parameter estimation accuracy. The proposed algorithm can be extended to other nonlinear systems (Yu, Mao, Jia, & Yuan, 2014) and time-delay systems (Li & Shi, 2012).

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0.17129

0.48730

31.52362

 $\gamma_1$ 

 $\gamma_2 \\ \delta(\%)$ 

0.08632

0.29787

16.88089

t	100	200	500	1000	2000	3000	True
$a_1$	0.51359	0.53236	0.55166	0.56302	0.56413	0.57026	0.84000
$a_2$	0.29217	0.30890	0.32556	0.33592	0.33676	0.34249	0.13000
$b_1$	0.64088	0.64163	0.64053	0.64003	0.63957	0.63947	0.71000
$b_2$	0.76764	0.76701	0.76794	0.76835	0.76873	0.76882	0.70420
$h_1$	1.04698	1.02740	0.93650	0.93191	0.92790	0.92554	0.43000
$\gamma_1$	0.01925	0.05292	0.03502	0.03106	0.03033	0.02763	0.05000
$\gamma_2$	0.45292	0.41385	0.37070	0.33491	0.29097	0.27826	0.09000
$\delta(\%)$	58.43671	55.79285	49.28239	47.91747	46.60821	46.15376	
Table 2. The MISG estimates and errors with $p = 3$							
t	100	200	500	1000	2000	3000	True
$a_1$	0.66639	0.68320	0.71943	0.72460	0.72949	0.73495	0.84000
$a_2$	0.17307	0.18519	0.21861	0.22351	0.22822	0.23356	0.13000
$b_1$	0.66106	0.65973	0.66033	0.66049	0.66048	0.66050	0.71000
$b_2$	0.75033	0.75150	0.75098	0.75083	0.75084	0.75083	0.70420
$h_1$	0.74457	0.74056	0.73754	0.73651	0.73403	0.73225	0.43000
$\gamma_1$	0.12732	0.11880	0.05048	0.02801	0.01935	0.01763	0.05000
$\gamma_2$	0.45465	0.34179	0.24566	0.16457	0.11551	0.09950	0.09000
$\delta(\%)$	37.82532	32.04669	27.57636	25.69972	25.06640	24.89608	
Table 3. The MISG estimates and errors with $p = 8$							
t	100	200	500	1000	2000	3000	True
$a_1$	0.76889	0.78661	0.83522	0.84053	0.84585	0.84883	0.84000
$a_2$	0.07563	0.08015	0.12405	0.12874	0.13381	0.13672	0.13000
$b_1$	0.71084	0.70931	0.70940	0.70937	0.70935	0.70934	0.71000
$b_2$	0.70336	0.70490	0.70480	0.70484	0.70486	0.70486	0.70420
$h_1$	0.32987	0.36079	0.39332	0.40920	0.42416	0.43059	0.43000

0.03428

0.07007

2.37246

0.02372

0.07667

2.22821

0.03227

0.07139

2.02323

0.05000

0.09000

Table 1. The MISG estimates and errors with p = 1



0.02044

0.14186

5.08940

Fig. 1. Parameter estimation errors  $\delta$  versus t with different innovation length p

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